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# Application of Biomarkers in Cancer Epidemiology

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## General issues of study design and analysis in the use of biomarkers in cancer epidemiology

N. Pearce and P. Boffetta

Other contributions to this volume have discussed sources of variation (see Vineis) and measurement error (see White). In this article, we focus on statistical issues involved in the design and analysis of epidemiological studies that use biomarkers. We do not consider statistical issues of laboratory analyses.

In general, epidemiological research involves studying the external, modifiable causes of diseases in populations (McMichael, 1994, 1995) with the intention of developing preventive interventions. In some instances, this activity can be enhanced by using internal biomarkers to obtain better measurements of internal exposure (dose), disease or individual susceptibility. Statistical issues in studies using biomarkers of exposure are not markedly different from those involved in other epidemiological studies based on measures of external exposure, while studies using biomarkers of disease pose specific problems due to the lack of persistence of some of these markers, and the analysis of interaction is of particular interest in studies using markers of susceptibility. As with other epidemiological studies, the statistical analysis of a study involving biomarkers involves in general: (1) relating a particular disease (or health outcome, such as a marker of early effect) to (2) a particular exposure while (3) controlling for systematic error, (4) assessing interactions with other exposures and (5) assessing the possibility of random error. We will consider each of these five aspects of study design and analysis in turn. We will restrict our discussion to full-scale epidemiological studies, i.e. we assume that the use of biomarkers, be they of exposure, effect or susceptibility, aims to contribute to the elucidation of the causal relationships in human populations between diseases and factors such as external exposures, personal habits, genetic traits and interventions. Issues in the design and analysis of transitional studies, in which the main aim is the validation of the markers themselves, are outside the scope of this cha əter.

### Measuring disease with biomarkers

Epidemiological studies are usually based on a particular population followed over a particular period of time. Miettinen (1985) has termed this study population the 'base population' and its experience over time the 'study base'. The different epidemiological study designs differ only in the manner in which the study base is defined and the manner in which information is drawn from the study base (Checkoway et al., 1989). Thus, epidemiological studies may involve measuring either incidence or prevalence of disease. This distinction is important when a biomarker is being used to measure either the disease under study or early biological effects that are considered to be valid predictors of disease risk (e.g. Rothman et al., 1995). In particular, many studies measuring disease with biomarkers are of cross-sectional design and measure the prevalence of the disease, which is dependent on its incidence and its duration. Thus, in a study looking at markers of cell damage as an effect of exposure to known or suspected carcinogens, the results would depend on factors such as the turnover of the cells in which the marker is measured or the capacity to repair the damage. For example, there is evidence that chromosomal aberrations caused by some carcinogens, such as arsenic and benzene, last for longer periods than aberrations caused by vinyl chloride, and the reason for this difference is not known (Schwartz, 1990).

### Incidence studies

Three measures of disease incidence are commonly used in incidence studies. The (person-time) incidence rate (or incidence density; Miettinen,

1985) is a measure of the disease occurrence per unit time. A second measure of disease occurrence is the cumulative incidence (incidence proportion; Miettinen, 1985) or risk, which is the proportion of study subjects who experience the outcome of interest at any time during the follow-up period. A third possible measure of disease occurrence is the incidence odds (Greenland, 1987), which is the ratio of the number of subjects who experience the outcome to the number who do not experience the outcome. As for the cumulative incidence, the incidence odds is dimensionless, but it is necessary to specify the time period over which it is being measured. In incidence studies involving biomarkers of disease, it is therefore important to consider whether a particular biomarker measures incidence or cumulative incidence.

Corresponding to these three measures of disease occurrence, there are three principal ratio measures of effect that can be used in cohort studies. The measure of primary interest is often the rate ratio (incidence density ratio), which is the ratio of the incidence rate in the exposed group to that in the non-exposed group. A second commonly used effect measure is the risk ratio (cumulative incidence ratio), which is the ratio of the cumulative incidence in the exposed group to that in the non-exposed group. When the outcome is rare over the follow-up period, the risk ratio is approximately equal to the rate ratio. A third possible effect measure is the incidence odds ratio, which is the ratio of the incidence odds in the exposed group to that in the non-exposed group. An analogous approach can be used to calculate measures of effect based on the differences rather than the ratios, in particular the rate difference and the risk difference.

In incidence studies involving biomarkers of disease, it is therefore important to consider whether a particular biomarker measures incidence or cumulative incidence. For example, if the 'disease' under study is hepatitis B virus (HBV) infection, then a survey of the prevalence of HBV markers, in a cohort that has been followed over time, will indicate the cumulative incidence of infection in the cohort (with the exception of those who have died from any cause during follow-up or who no longer show evidence of infection). It will not directly indicate the incidence rate of infections; this would require repeated prevalence surveys over time.

Incidence case-control studies

Incidence case-control studies involve studying all of the incident cases of disease generated by the study base and a control group sampled at random from the same study base. The relative risk measure is the incidence odds ratio; the effect measure that this estimates depends on the manner in which controls are selected. Once again, there are three main options (Pearce, 1993).

One option is to select controls from those who do not experience the outcome during the followup period, i.e. the survivors (those who did not develop the outcome at any time during the followup period). In this instance, a sample of controls chosen by cumulative incidence sampling will estimate the exposure odds of the survivors, and the odds ratio obtained in the case-control study will therefore estimate the incidence odds ratio in the base population. Controls can also be sampled from the entire base population (those at risk at the beginning of follow-up), rather than just from the survivors (those at risk at the end of follow-up). In such case-base sampling, the controls will estimate the exposure odds in the base population of persons at risk at the start of follow-up, and the odds ratio obtained in the case-control study will therefore estimate the risk ratio in the base population. The third approach is to select controls longitudinally throughout the course of the study (Miettinen, 1976); this is sometimes described as 'risk-set sampling' (Robins et al., 1986), 'sampling from the study base' (the person-time experience; Miettinen, 1985) or 'density sampling' (Kleinbaum et al., 1982). In this instance, the controls will estimate the exposure odds in the study base, and the odds ratio obtained in the case-control study will therefore estimate the rate ratio in the study base.

In incidence case—control studies involving biomarkers of disease, it is therefore important to consider whether a particular biomarker measures incidence or cumulative incidence. These issues determine not only which measure of effect is being estimated, but also which method of control selection is appropriate, and the resulting methods of data analysis.

### Prevalence studies

The term prevalence denotes the number of cases of disease existing in the population at the time the study was conducted. If we denote the prevalence

of a disease in the study passume that the inciden time, the population has and there is no migration lence pool, then it can be that the prevalence odds rate (I) multiplied by the (D):

$$P/(1-P) = I \times D$$

Thus, the prevalence ational to the disease incide odds ratio is estimated to

$$OR = I_1 D_1 / I_0 D_0$$

An increased prevalence flect the influence of factors ease incidence and/or fact duration. The different unincreasing disease incident likely to involve different sure and disease (see belikely to require different unent of the etiologically in the control of the etiologically in the control of the etiologically in the et

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es the number of cases ulation at the time the denote the prevalence

of a disease in the study population by P, and if we assume that the incidence rate is constant over time, the population has reached a 'steady state' and there is no migration into or out of the prevalence pool, then it can be shown (Rothman, 1986) that the prevalence odds is equal to the incidence rate (I) multiplied by the average disease duration (D):

$$P/(1-P) = I \times D$$

Thus, the prevalence odds is directly proportional to the disease incidence, and the prevalence odds ratio is estimated to be:

$$OR = I_1 D_1 / I_0 D_0$$

An increased prevalence odds ratio may thus reflect the influence of factors that increase the disease incidence and/or factors that increase disease duration. The different mechanisms involved in increasing disease incidence or disease duration are likely to involve different time patterns of exposure and disease (see below), which in turn are likely to require different biomarkers for measurement of the etiologically relevant exposures.

### Prevalence case-control studies

Just as an incidence case-control study can be used to obtain the same findings as a full cohort study, a prevalence case-control study can be used to obtain the same findings as a full prevalence study in a more efficient manner. In particular, if obtaining exposure information is difficult or costly (e.g. if it involves serum samples), then it may be more efficient to conduct a prevalence case-control study by obtaining exposure information on all of the prevalent cases of the disease under study and a sample of controls selected at random from the non-cases. In this instance, a sample of controls will estimate the exposure odds of the non-cases, and the odds ratio obtained in the prevalence case-control study will therefore estimate the prevalence odds ratio in the base population, which in turn estimates the incidence rate ratio, provided that the average duration of disease is the same in the exposed and non-exposed groups. Once again, an increased prevalence odds ratio may reflect the influence of factors that increase disease incidence and/or factors that increase disease duration, and

the different mechanisms involved are likely to require different biomarkers for measurement of the etiologically relevant exposures.

### Measuring exposure with biomarkers

Validity of biomarkers of exposure

There are considerable shortcomings in many currently available biomarkers of exposure, including problems of measuring historical exposures; uncertainties as to what a biomarker is measuring: greater susceptibility to confounding in some instances; problems of application to public health policy (Pearce et al., 1995); the disease process affecting the level of the biomarker; and problems of validity of laboratory measurements (Boffetta. 1995). These issues are covered elsewhere in this volume. In this section we concentrate on issues of study design and analysis when measuring exposure with biomarkers, particularly with regard to time-related exposures with a relatively long induction and latency period between exposure and the subsequent occurrence of disease (as is the situation in most cancer epidemiology studies). The issues we discuss are not unique to a particular study design (cohort studies, case-control studies, cross-sectional studies) but rather apply to all studies in which the etiologically relevant time period involves a relatively long induction and latency period, thereby posing problems with the measurement of exposure during this period.

### Time-related exposures

Some biomarkers measure factors that are fixed and do not change over time in an individual, e.g. genetic susceptibility genes that may interact with xenobiotic factors in cancer causation (Rothman, 1995). Other biomarkers measure factors that change over time, e.g. micronutrient levels in serum may change from day-to-day (Willett, 1990).

In studies (both prospective and retrospective) of long-term health effects involving time-related exposures, it is important that the time patterns of the study exposure and of the relevant confounders should be taken into account in the analysis (Pearce et al., 1986). In particular, it is important that the principal exposure under study should be analysed in a time-related manner, taking account of the likely induction and latency periods, and the relative etiological importance of exposure intensity, exposure duration and

cumulative exposure. The simplest approach is to analyse the cumulative exposure in a time-related manner, and this may suffice when the aim is merely to consider whether or not there is an effect of exposure. However, once it has been provisionally assumed that an effect exists, attention then shifts to understanding the nature of the effect. In this context, the temporal pattern of exposure and outcome can be considered by examining the effects of exposures in specific time windows while controlling for time-related confounders and for the effects of exposures in other time windows. A more sophisticated approach is direct fitting of a theoretical model of carcinogenesis (Pearce, 1992), which requires assumptions as to the relevance of the times in which the exposure occurred and in which the marker was measured. This may be particularly relevant in studies including the measurement of the exposure, the marker and the disease, which are mainly aimed at elucidating the role of the marker in the exposure-disease relationship (Schatzkin et al., 1993). An example of this type of study is the investigation of the role of human papilloma virus (HPV) in the etiology of cervical cancer (the 'exposure' in this case being factors such as number of sexual partners and age at first intercourse). In this case, the association with HPV infection, when properly measured with PCR-based assays, is of such a magnitude that there is little concern about the relevance of the time periods (Muñoz et al., 1992). However, for cancers from other organs, the association with HPV infection is less clear-cut, and the relevance of the timing of the infection may be one of the unknown factors modifying the exposure-marker-disease relationship.

Thus, biological measures of time-related exposures must be able to measures changes in exposure levels over time. In particular, stored biological samples may not provide valid measurements of long-term patterns of exposure when there are significant variations in exposure over time, unless samples have been taken repeatedly over the course of the study (Armstrong et al., 1992). If it is not possible to take repeated biological samples over time, then it is essential that the samples that are taken relate to the etiologically relevant time period.

Many currently available biomarkers only indicate relatively recent exposures. For example, it is well known that serum levels of micronutrients re-

flect recent rather than historical dietary intake (Willett, 1990); given the long induction time of most cancers it is usually exposures between 10 and 30 years previously that are etiologically relevant. While this may not be a limitation in crosssectional studies, provided that the etiologically relevant time period is close to the time of data collection, it is an important limitation in cohort and case-control studies aiming to assess the effects of historical exposures. Some biomarkers are better than others in this respect (particularly biomarkers for exposure to biological agents), but even the best markers of chemical exposures reflect only the last few weeks or months of exposure. On the other hand, with some biomarkers (e.g. serum levels of TCDD; Johnson et al., 1992) it may be possible to estimate historical levels if the exposure period is known, if the half-life is relatively long (and is known) and if it is assumed that no significant exposure has occurred more recently, or if it is reasonable to assume that exposure levels have remained stable over time.

However, historical information on exposure surrogates will often be more valid than current direct measurements of exposure or dose. This situation has long been recognized in occupational epidemiology, where the use of work history records in combination with a job-exposure matrix (based on historical exposure measurements of work areas rather than individuals) is usually more valid than current exposure measurements (whether based on environmental measurements or biomarkers) because of changes in exposure levels over time (Checkoway et al., 1989). Similar problems may occur in the measurement of other carcinogenic exposures. For example, even the best currently available measures of exposure to tobacco smoke, such as plasma or urinary cotinine, appear to have similar validity to questionnaires for the measurement of current exposures; their very short halflife makes them inferior to questionnaires in the estimation of historical exposures (Pearce et al., 1995). On the other hand, some biomarkers would appear to have value in the validation of questionnaires (Forastiere et al., 1993), which can then be used to estimate historical exposures.

Another example in which timing of sample collection is of great importance is in the case of DNA adducts (Wilcosky & Griffith, 1990). Since most adducts are readily repaired, any measure of

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### Systematic error

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### Selection bias

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timing of sample ice is in the case of iffith, 1990). Since ed, any measure of exposure based on DNA adducts will depend on the time between the end of exposure and sample collection; this pattern can then be modified by factors such as the activity of repair enzymes, which in turn may also have an independent influence on the outcome, i.e. may be associated with case—control status. DNA adduct formation and repair are particularly problematic, since:

- the extent to which the amount of adducts measured represents the amount of biologically active adducts formed (as discussed above) is usually not known; and
- in the case of measurements taken during exposure (or immediately thereafter, such as at the end of a working day) it is not known how much of the adducts found would persist long enough to be biologically important.

### Analysis based on pooled samples

These issues of the timing of sample collection are of particular concern in analyses based on pooled samples. If the samples were not all taken at the same etiologically relevant time period, then the pooled sample will represent the average of samples taken from a variety of time periods. Thus, there may be considerable misclassification with regard to the exposure levels at the etiologically relevant time period.

### Systematic error

The major possible types of systematic error (bias) are the same in traditional epidemiology and in studies involving biomarkers (Boffetta, 1995). The various types of bias can be grouped into three major classes: selection bias, information bias, and confounding (Rothman, 1986). This section is not intended to give a comprehensive review of these types of bias; rather, we will concentrate on issues involving data analysis. One solution is to pool small sets of samples stratified on the basis of time since collection; however, this procedure may substantially reduce the advantages of pooling.

### Selection bias

Selection bias involves biases arising from the procedures by which the study participants are chosen from the study base. Selection bias can be avoided by including all of the study base (i.e. a cohort study) and obtaining a response rate of 100%. This

is often not practicable, but selection bias can also be controlled in the analysis by identifying factors that are related to subject selection and controlling for them as confounders. The statistical issues involved in controlling for selection bias in the analysis are essentially the same as those involved in controlling for sources of confounding (see below).

### Information bias

Information bias involves misclassification of the study participants with respect to disease or exposure status. Thus, the concept of information bias refers to those people actually included in the study, whereas selection bias refers to the selection of the study participants from the study base, and confounding generally refers to non-comparability of subgroups within the study base. The various methodological issues of validity, reproducibility and stability of markers are part of the more general problem of information bias.

Non-differential information bias occurs when the likelihood of misclassification of exposure is the same for both cases and non-cases of disease (or when the likelihood of misclassification of disease is the same for exposed and non-exposed persons). Non-differential misclassification of exposure generally biases the relative risk estimate towards the null value of 1.0 (Copeland et al., 1977). Hence, non-differential information bias tends to produce 'false negative' findings and is of particular concern in studies that find no association between exposure and disease.

Differential information bias occurs when the likelihood of misclassification of exposure is different between cases and non-cases (or the likelihood of misclassification of disease is different between exposed and non-exposed persons). This can bias the observed effect estimate in either direction, either towards or away from the null value.

Information bias can drastically affect the validity of a study. As a general principle, it is important to ensure that the misclassification is non-differential, by ensuring that exposure information is collected in an identical manner in cases and non-cases (or that disease information is collected in an identical manner in the exposed and non-exposed groups). In this situation, the bias is in a known direction (towards the null), and although there may be concern that not finding a significant association

(between exposure and disease) may be due to non-differential information bias, at least one can be confident that any positive findings are not due to information bias. Thus, the aim of data collection is not to collect perfect information, but to collect comparable information in a similar manner from the groups being compared, even if this means ignoring more detailed exposure information if this is not available for both groups. However, it is clearly important to collect information that is as detailed and accurate as possible, within the constraints imposed by the need to ensure that information is collected in a similar manner in the groups being compared.

In general, cross-sectional and case-control studies based on biomarkers of exposure are more prone to differential misclassification than studies based on measurement of external exposure and disease, since the biomarkers may be influenced by the disease itself. This problem is less relevant in prospective studies (or nested case-control studies) in which the marker is measured on biological material collected before the onset of disease, provided that cases diagnosed within a short interval after sample collection are excluded. The fact that the relationships between exposure, marker and disease are in most cases obscure limits the interpretation of the findings of biomarker-based studies with respect to the presence or absence of information bias.

### Confounding

Confounding occurs when the exposed and nonexposed groups (in the study base) are not comparable due to inherent differences in background disease risk (Greenland & Robins, 1986) caused by exposure to other risk factors. The concept of confounding thus generally refers to the study base, although as noted above, confounding can also be introduced (or removed) by the manner in which study participants are selected from the study base.

If no other biases are present, three conditions are necessary for a factor to be a confounder (Rothman, 1986). First, a confounder is a factor that is predictive of disease (in the absence of the exposure under study); second, a confounder is associated with exposure in the study base; and third, a variable that is intermediate in the causal pathway between exposure and disease is not a confounder. This latter issue is of particular con-

cern in studies using biomarkers, since the identification of potential confounders depends on previous knowledge of the relationship between these and the relevant variables of exposure and outcome, and such knowledge is, for most biomarkers, very poor.

The most straightforward method of controlling confounding in the analysis involves stratifying the data into subgroups according to the levels of the confounder(s) and calculating a summary effect estimate that summarizes the information across strata. However, it is usually not possible to control simultaneously for more than two or three confounders when using stratified analysis. This problem can be mitigated to some extent by the use of mathematical modelling, but this may in turn produce problems of multicollinearity when variables which are highly correlated are entered simultaneously into the model. For example, serum levels of various micronutrients may be strongly correlated, and multicollinearity may occur when they are included in the same model. This will lead to unstable effect estimates with large standard errors, and may in fact lead to the 'wrong' micronutrient showing the strongest association (negative or positive) with disease. This may be one reason why numerous studies have shown that the consumption of green and yellow vegetables is protective against a range of cancers, but the identification of the specific dietary micronutrients involved has remained elusive (Steinmetz & Potter, 1991).

In general, control of confounding requires careful use of *a priori* knowledge, together with assessment of the extent to which the effect estimate changes when the factor is controlled in the analysis. Most epidemiologists prefer to make a decision based on the latter criterion, although it can be misleading, particularly if misclassification is occurring (Greenland & Robins, 1985). The decision to control for a presumed confounder can certainly be made with more confidence if there is supporting prior knowledge that the factor is predictive of disease.

Misclassification of a confounder leads to a loss of ability to control confounding, although control may still be useful provided that misclassification of the confounder was unbiased (Greenland, 1980). Misclassification of exposure is more problematic, since factors that influence misclassifica-

tion may appear to I these factors may in & Robins, 1985).

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When appropriate information is not available to control confounding directly, it is still desirable to assess its potential strength and direction. For example, it may be possible to obtain information on a surrogate for the confounder of interest. Even though confounder control will be imperfect in this situation, it is still possible to examine whether the main effect estimate changes when the surrogate is controlled in the analysis, and to assess the strength and direction of the change. Alternatively, it may be possible to obtain accurate confounder information for a subgroup of participants (cases and non-cases) in the study and to assess the effects of confounder control in this subgroup.

A related approach involves obtaining confounder information for a sample of the study base (or a sample of the controls in a case—control study). For example, in a study based on questionnaires, biomarkers may be used to validate questionnaire information in a subgroup of study participants.

The potential for confounding is of major concern in all epidemiological studies, including those using biomarkers. The use of biomarkers of exposure does not reduce the need to control for confounding, and in some instances the use of biomarkers may actually introduce confounding into a study. For example, in a study of lung cancer and PAH exposure in a group of factory workers, if the workers are classified according to PAH exposure on the basis of industrial hygiene monitoring, then the percentage of smokers (and the mean number of cigarettes smoked per day) will usually be similar in the groups with low, medium and high levels of occupational exposure to PAH (since these groups are defined purely on environmental levels of PAH exposure in various job categories and departments, and these exposures will usually be unrelated to cigarette smoking). In this situation, cigarette smoking will not be a major confounder. However, if workers are classified according to PAH-DNA levels, these will indicate total exposure to PAHs from all sources, including cigarette smoking. Thus, cigarette smokers will be more likely to be in the 'high PAH exposure' group, and this group will therefore contain a higher proportion of smokers (and a higher mean number of cigarettes smoked

per day) than the medium or low PAH exposure group (since some of the total PAH exposure comes from cigarette smoke). The dose-response will then be confounded by cigarette smoking, and the 'high PAH exposure' group may show a higher lung cancer risk which is not due to PAII exposure, but which is actually due to the other carcinogenic constituents of cigarette smoke (Pearce et al., 1995). One solution is to stratify the analysis on cigarette smoking (as measured by questionnaire), but any confounding control is likely to be imperfect. This is even more of a problem if biomarkers are used to measure tobacco smoking because of the problems of measuring the etiologically relevant constituents of tobacco smoke (as distinct from exposure to PAHs in tobacco smoke). Once again, confounding control is likely to be imperfect and, therefore, to yield results that are still confounded and less valid than those obtained by only considering occupational exposures (using a job-exposure matrix). Furthermore, one can only control for known confounders (e.g. tobacco smoke) and cannot control for unknown confounders that may also be subject to the same types of biases described above. Therefore, it is usually preferable to avoid confounding rather than to attempt to control for it post hoc (which is why randomized trials are the preferred method when they are feasible). Thus, it is preferable to consider only occupational exposures to PAHs, using a job-exposure matrix, and not to attempt to measure non-occupational exposures to PAH using biomarkers. Finally, it should be stressed that the above issues have been discussed in terms of studies in which exposure is measured prospectively; the problems are even more acute when historical exposures are being assessed.

### Random error

Random error will occur in any epidemiological study, just as it occurs in experimental studies. It is often referred to as 'chance', although it can perhaps more reasonably be regarded as 'ignorance' (Checkoway et al., 1989). Even in an experimental study in which participants are randomized into 'exposed' and 'non-exposed' groups, there will be 'random' differences in background risk between the compared groups, but these will diminish in importance (i.e. the random differences will 'even out') as the study size grows. In epidemiological

studies, there is no guarantee that differences in baseline (background) risk will even out between the exposure groups as the study size grows, but it is necessary to make this assumption in order to proceed with the study (Greenland & Robins, 1986). In practice, the study size depends on the number of available participants and the available resources. Within these limitations, it is desirable to make the study as large as possible, taking into account the trade-off between including more participants and gathering more detailed information about a smaller number of participants (Greenland, 1988).

A major problem with the use of biomarkers of exposure and outcome in cancer epidemiology studies (and particularly in cohort studies) is that of small numbers. Even large multicentre cohort studies often struggle to obtain sufficient numbers to assess risks of rare cancers from occupational exposures (Fingerhut et al., 1991; Saracci et al., 1991). The use of biomarkers may be a major problem in this regard, since the resulting expense and complexity may drastically reduce the study size, even in community-based and nested case-control studies, and therefore greatly reduce the statistical power for detecting an association between exposure and disease. As in traditional epidemiology studies, in studies using biomarkers statistical power is a function of the prevalence of exposure and the magnitude of risk; a biomarker with low prevalence and high relative risk can be evaluated in small populations, whereas a biomarker with low prevalence and low relative risk requires a larger population. The optimal balance between precision and validity depends on a number of considerations, including the relative costs of the various exposure measurement techniques (Greenland, 1988). Thus, for an expensive biomarker to be useful it must be substantially better than less expensive and less invasive approaches. However, it has been argued that the necessary study size in some molecular epidemiology studies may be smaller than in traditional epidemiology (Hulka, 1990a,b; Hertzberg & Russek-Cohen, 1993) because of larger differences in biomarker distribution, identification of subgroups at higher risk, and the use of continuous outcome variables (Boffetta, 1995). While this is true in many cases (e.g. the detection of mutations in critical genes as a marker of increased risk of cancer), in other cases the biomarkers may show a

very low (or very high) prevalence, thus requiring large samples to detect a difference between groups (Hulka & Margolin, 1992; Rothman et al., 1995). Finally, it should be noted that some biomarkers are of interest in themselves, rather than functioning as surrogates for other exposures; in particular, no alternative methods exist for measuring markers of genetic susceptibility. Nevertheless, the additional information provided by the use of such markers should still be compared with that provided by alternative, larger studies in which the marker is not used.

An additional consideration in study size estimation in studies using biomarkers is the ratio of the number of assays per individual and the number of individuals in the study (Boffetta, 1995), such as the detection of chromosomal aberrations or sister chromatid exchanges. In this case, studies must be based on adequate numbers of individuals and observations per individual (Hirsch et al., 1984; Whorton, 1985). Many biomarkers show marked variation from day to day within the same individual, and the intra-individual variation may be greater than the interindividual variation (Armstrong et al., 1992). It may therefore be necessary to take a large number of measurements to accurately estimate the average exposure level for each individual; otherwise it may be impossible to detect differences between individuals. For example, for 24-hour urinary sodium, the within-person variation may be three times as high as the between-person variation; it has been estimated that the misclassification resulting from taking only one sample per person would result in a true relative risk of 2.0 being reduced to an observed relative risk of 1.2 (Armstrong et al., 1992). Thus, it may be necessary to take 10-15 24-hour urine samples in order to achieve reasonable accuracy in estimating average individual sodium intake levels.

### Interaction

Interaction (effect modification) occurs when the estimate of effect of exposure depends on the level of another factor in the study base (Miettinen, 1974). The term statistical interaction denotes a similar phenomenon in the observed data. The former term will generally be used here. Interaction is distinct from confounding (or selection or information bias) in that it does not represent a bias which should be removed or controlled, but rather

a real difference in the c subgroups that may be example, in a cohort str asthma in children. passive smoking might groups, or in males: example of interaction hazardous in one group group. More generally. in both groups, but the vary. A typical exampl studies using biomarke of disease due to an ext the population with a bility marker, such as th zyme implicated in the of the agent (see Land In this situation, effect terpreted with conside ence of statistical interstatistical methods usc factors modify either t ference, as uniformity non-uniformity over ti the joint effects of tw goal of the study, this stratum-specific effect to proceed if statistical assessment of joint eff Some authors (e.g. Kk that it is not approprize late an overall estimate levels of the effect me mon to ignore this sti effect estimates is not fact, valid analytical rate ratios) have been this situation (Rothm?

Biomarkers may perfor assessing interact genetic and/or enviro lar, genetic susceptible higher disease risk for than for non-susceptil (Boffetta, 1995).

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n) occurs when the lepends on the level ly base (Miettinen, teraction denotes a served data. The forthere. Interaction is selection or inforot represent a bias introlled, but rather

a real difference in the effect of exposure in various subgroups that may be of considerable interest. For example, in a cohort study of passive smoking and asthma in children, the effect estimate for passive smoking might be different in different age groups, or in males and females. The clearest example of interaction is when a factor is actually hazardous in one group and protective in another group. More generally, the risk might be elevated in both groups, but the strength of the effect may vary. A typical example of effect modification in studies using biomarkers is the estimate of the risk of disease due to an external agent in subgroups of the population with a different genetic susceptibility marker, such as the polymorphism for an enzyme implicated in the activation or detoxification of the agent (see Landi & Caporasi, this volume). In this situation, effect modification should be interpreted with considerable care, since the presence of statistical interaction may depend on the statistical methods used. In fact, all secondary risk factors modify either the rate ratio or the rate difference, as uniformity over one measure implies non-uniformity over the other. If the assessment of the joint effects of two factors is a fundamental goal of the study, this can be done by calculating stratum-specific effect estimates. It is less clear how to proceed if statistical interaction is occurring, but assessment of joint effects is not an analytical goal. Some authors (e.g. Kleinbaum et al., 1982) argue that it is not appropriate in this situation to calculate an overall estimate of effect summarized across levels of the effect modifier. However, it is common to ignore this stipulation if the difference in effect estimates is not too great (Pearce, 1989). In fact, valid analytical methods (e.g. standardized rate ratios) have been specifically developed for this situation (Rothman, 1986).

Biomarkers may provide better opportunities for assessing interactions between two or more genetic and/or environmental factors. In particular, genetic susceptibility genes should produce a higher disease risk for exposed susceptible groups than for non-susceptible and non-exposed groups (Boffetta, 1995).

However, a major problem of testing for interaction, e.g. in studies involving markers of genetic susceptibility, is that it usually requires a substantial increase in study size. For example, in a case—control study, testing for interaction involves com-

paring the sizes of the odds ratios (relating exposure and disease) in different strata of the effect modifier, rather than merely testing whether the overall odds ratio is different from the null value of 1.0. The power of the test for interaction therefore depends on the numbers of cases and controls in specific strata (of the effect modifier) rather than the overall numbers of cases and controls. For example, Smith and Day (1984) give an example of a case-control study that would have to be five times larger to detect a difference between odds ratios of 1.0 and 2.0 in the two different strata of an effect modifier than it would have to be to detect an overall odds ratio of 2.0 (ignoring the effect modifier). In general, when considering possible interactions, the size of the study needs to be at least four times larger than when interaction is not considered (Smith & Day, 1984). Thus, in a study involving markers of genetic susceptibility, the gain in statistical power from considering such markers (thereby yielding higher relative risks in some strata) may be offset by the decrease in statistical power from the need to consider interactions. However, if the exposure of interest is independent from the genetic factor under study, case-case comparisons can be used to study interactions with greater statistical power (Piegorsch et al., 1984).

### Conclusions

In some instances, the increasing use of biomarkers in epidemiological studies represents a major improvement in the discipline during the last years (Vineis, 1992). In many cases, biomarkers have helped to improve our knowledge of causes and mechanisms of both disease etiology and prevention. In other cases, however, it is unclear whether they represent an improvement on traditional epidemiological methods. Epidemiological studies based on biomarkers are usually more complex than traditional epidemiological studies, because information is available on a larger number of variables whose biological meaning is often poorly known. The methodological considerations involved in classical epidemiological studies on issues such as measurement of disease, measurement of exposure, selection bias, confounding, precision and interaction also apply to biomarkerbased studies, and in most cases the methodological problems of this type of study do not require solutions different from those used in classical studies. In some cases, however, the use of biomarkers may pose specific problems, which have to be addressed within the general framework of modern epidemiological methods.

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